

## Lecture 3/

### Chapter 7 Properties of Expectation.

- this chapter focuses on developing and exploring further properties of expected value.

Prop: Suppose  $X$  is a RV s.t.  $P(a \leq X \leq b) = 1$ .  
Then  $a \leq E[X] \leq b$ .

Pf (Continuous case):

$$\int_a^b x f(x) dx \leq E[X] \leq \int_a^b b f(x) dx$$

$\Downarrow$

$$a \int_a^b f(x) dx \leq E[X] \leq b \int_a^b f(x) dx$$

$a \leq E[X] \leq b$

*Handwritten notes:*  
- Red arrow from  $a$  to  $\int_a^b x f(x) dx$  with text "x bounded below by a"  
- Red arrow from  $b$  to  $\int_a^b b f(x) dx$  with text "x bounded above by b"

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### 7.2. Expectations of Sums of RV's

Prop: Suppose  $X$  and  $Y$  have:

- joint prob mass function  $p(x,y)$ , then

$$E[g(X, Y)] = \sum_x \sum_y g(x, y) p(x, y).$$

b) joint density  $f(x, y)$ , then

$$E[g(x, y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f(x, y) dx dy.$$

See the text for proof.

Special case:  $g(x, y) = ax + by$

Then

$$E[ax + by] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (ax + by) f(x, y) dx dy.$$

$$= a \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x f(x, y) dy dx + b \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} y f(x, y) dx dy$$

$$= a \int_{-\infty}^{\infty} x f_X(x) dx + b \int_{-\infty}^{\infty} y f_Y(y) dy.$$

$$= aE[X] + bE[Y], \text{ whenever } E[X] \text{ and } E[Y] \text{ are finite.}$$

An easy induction argument gives more generally that

$$E\left[c + \sum_{i=1}^n a_i X_i\right] = c + \sum_{i=1}^n a_i E[X_i].$$

### Example: (Sample mean)

Suppose  $X_1, \dots, X_n$  be independent, identically distributed (iid) with mean  $\mu$ . Then  $\bar{X} = \frac{\sum_{i=1}^n X_i}{n}$  is called the sample mean and  $E[\bar{X}] = \frac{1}{n} \sum_{i=1}^n E[X_i] = \frac{n\mu}{n} = \mu$ .

### Ex: unions of events.

Let  $A_1, A_2$  be events.

Defn  $X_i = \begin{cases} 1 & \text{if } A_i \text{ occurs} \\ 0 & \text{otherwise.} \end{cases}$

Observe that

$$\begin{aligned} 1 - (1 - X_1)(1 - X_2) &= 1 - (1 - X_1 - X_2 + X_1 X_2) \\ &= X_1 + X_2 - X_1 X_2 \\ &= \begin{cases} 1 & \text{if } A_1 \cup A_2 \text{ occurs} \\ 0 & \text{otherwise.} \end{cases} \end{aligned}$$

Now,  $E[X_i] = P(A_i)$ .

$$\text{So } E[1 - (1 - X_1)(1 - X_2)] = P(A_1 \cup A_2).$$

$$\begin{aligned} \text{S } E[X_1 + X_2 - X_1 X_2] &= P(A_1 \cup A_2) \\ E[X_1] + E[X_2] - E[X_1 X_2] &= P(A_1 \cup A_2). \end{aligned}$$

$$\Rightarrow P(A_1) + P(A_2) - P(A_1 A_2) = P(A_1 \cup A_2).$$

Same argument shows inclusion-exclusion for events  $A_1, \dots, A_n$ .

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Remark: In general, it is not true that

$$E\left[\sum_{i=1}^{\infty} X_i\right] = \sum_{i=1}^{\infty} E[X_i]. \quad (*)$$

why?  $E\left[\sum_{i=1}^{\infty} X_i\right] = E\left[\lim_{n \rightarrow \infty} \sum_{i=1}^n X_i\right]$

but the RHS is not necessarily

$$\lim_{n \rightarrow \infty} E\left[\sum_{i=1}^n X_i\right] \quad \text{i.e. can't swap "E" and "lim".}$$

Two cases where  $(*)$  IS true

1)  $X_i$  are all non-negative, i.e.  $P(X_i \geq 0) = 1$ .

2)  $\sum_{i=1}^{\infty} E[|X_i|] < \infty$ .

Proposition: Suppose that  $X$  and  $Y$  are independent.  
Then, for any functions  $g(x)$ ,  $h(y)$ ,

$$E[g(X)h(Y)] = E[g(X)]E[h(Y)]$$

Pf. (Continuous case). Let  $f(x, y)$  be the joint density.

$$\begin{aligned} E[g(X)h(Y)] &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x)h(y)f(x, y) dx dy \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x)h(y) f_X(x) f_Y(y) dx dy \\ &= \int_{-\infty}^{\infty} g(x) f_X(x) dx \int_{-\infty}^{\infty} h(y) f_Y(y) dy \\ &= E[g(X)] E[h(Y)]. \end{aligned}$$

Defn: For random variables  $X$  and  $Y$ ,  
we define the covariance of  $X$  and  $Y$ , as

$$\text{cov}(X, Y) = E[(X - E[X])(Y - E[Y])].$$

Note that

$$\begin{aligned} \text{cov}(X, Y) &= E[XY - Y E[X] - X E[Y] + E[X]E[Y]] \\ &= E[XY] - E[X]E[Y] \end{aligned}$$

It follows that if  $X$  and  $Y$  are independent,

$$\begin{aligned}\text{Then } \text{cov}(X, Y) &= E[XY] - E[X]E[Y] \\ &= E[X]E[Y] - E[X]E[Y] = 0.\end{aligned}$$

The converse is not true:

Ex: let  $X$  be uniform on  $(-1, 1)$ . let  $Y = X^2$ .

$$\begin{aligned}\text{Then } \text{Cov}(X, Y) &= E[XY] - E[X]E[Y] \\ &= E[X^3] - 0 \cdot E[Y].\end{aligned}$$

$$= E[X^3] = \int_{-1}^1 x^3 \cdot 1 dx = \frac{x^4}{4} \Big|_{-1}^1 = \frac{1}{4} - \frac{1}{4} = 0.$$

So  $\text{cov}(X, Y) = 0$ , but  $X$  and  $Y$  are clearly not independent.